

Linear Regression

Given n points $(x_1, y_1) \dots (x_n, y_n)$ and an assumed relation $y = f(x) + \epsilon, \epsilon \sim N(\mu, \sigma)$ we want to find a model $\tilde{y}_i = ax_i + b$ such that the residual squared error

$$\text{RSS}(a, b) = \sum_{i=1}^n (\tilde{y}_i - y_i)^2$$

is minimized.

RSS is a function of the line parameters a and b . To minimize it we set both partial derivatives to zero. (This could technically find a maximum – but it's reasonably clear this function has no maximum value because the error can always be increased.)

Take partial derivatives

$$\begin{aligned}\frac{\partial \text{RSS}}{\partial a} &= 2 \sum (\tilde{y}_i - y_i) \frac{\partial}{\partial a} (\tilde{y}_i - y_i) \\ &= 2 \sum (\tilde{y}_i - y_i) (x_i) \\ \frac{\partial \text{RSS}}{\partial b} &= 2 \sum (\tilde{y}_i - y_i) \frac{\partial}{\partial b} (\tilde{y}_i - y_i) \\ &= 2 \sum (\tilde{y}_i - y_i) (1)\end{aligned}$$

Since

$$\frac{\partial}{\partial a} \tilde{y}_i = \frac{\partial}{\partial a} (ax_i + b) = x_i$$

$$\frac{\partial}{\partial b} \tilde{y}_i = \frac{\partial}{\partial b} (ax_i + b) = 1$$

And solve

$$\begin{cases} \frac{\partial \text{RSS}}{\partial a} = 0 \\ \frac{\partial \text{RSS}}{\partial b} = 0 \end{cases} \Rightarrow \begin{cases} \sum (\tilde{y}_i - y_i) x_i = 0 \\ \sum (\tilde{y}_i - y_i) = 0 \end{cases}$$

Since $\tilde{y}_i = ax_i + b$

$$\sum (ax_i + b - y_i) x_i = 0 \Rightarrow a \sum x_i^2 + b \sum x_i = \sum y_i x_i$$

and

$$\sum (ax_i + b - y_i) = 0 \Rightarrow a \sum x_i + b \sum 1 = \sum y_i$$

by Cramer's rule

$$a = \frac{\begin{vmatrix} \sum x_i y_i & \sum x_i \\ \sum y_i & n \end{vmatrix}}{\begin{vmatrix} \sum x_i^2 & \sum x_i \\ \sum x_i & n \end{vmatrix}}$$

$$b = \frac{\begin{vmatrix} \sum x_i^2 & \sum x_i y_i \\ \sum x_i & \sum y_i \end{vmatrix}}{\begin{vmatrix} \sum x_i^2 & \sum x_i \\ \sum x_i & n \end{vmatrix}}$$

since $\sum_{i=1}^n 1 = n$

Taking determinants,

$$a = \frac{n \sum x_i y_i - (\sum x_i)(\sum y_i)}{n \sum x_i^2 - (\sum x_i)^2}$$

$$b = \frac{\sum y_i \sum x_i^2 - \sum x_i \sum x_i y_i}{n \sum x_i^2 - (\sum x_i)^2}$$

Interpretation as a ratio of variances

Students of statistics may appreciate the following manipulations

Definition of covariance

$$E(xy) - E(x)E(y) = \text{Cov}(x, y)$$

Definition of variance

$$\text{Var}(x) = E[(x - \mu)^2]$$

Lemma

$$\begin{aligned} \text{Var}(x) &= E[(x - \mu)^2] \\ &= E(x^2) - 2\mu E[x] + E[\mu]^2 \\ &= E[x^2] - 2E[x]^2 + \mu^2 \\ &= E[x^2] - E[x]^2 \end{aligned}$$

Manipulating the denominator of the equation for a on the previous page,

$$\begin{aligned} n \sum x_i^2 - \left(\sum x_i\right)^2 &= n^2 \left(\frac{1}{n} \sum x_i^2 - \left(\frac{\sum x_i}{n}\right)^2 \right) \\ &= n^2 (E[x^2] - E[x]^2) \\ &= n^2 \text{Var}(x) \end{aligned}$$

And the numerator

$$\begin{aligned} n \sum x_i y_i - \sum x_i \sum y_i &= n^2 \left(\frac{1}{n} \sum x_i y_i - \frac{1}{n} \sum x_i \cdot \frac{1}{n} \sum y_i \right) \\ &= n^2 (E[xy] - E[x]E[y]) \\ &= n^2 (E[xy] - \mu_x \mu_y) \\ &= n^2 \text{Cov}(x, y) \end{aligned}$$

so

$$a = \frac{E[xy] - \mu_x \mu_y}{E[x] - \mu_x^2} = \frac{\text{Cov}(x, y)}{\text{Var}(x)}$$